Approximate Inference for partly observed

Continuous Time Markov Processes

Manfred Opper



Computer Science

collaboration with:

Andreas Ruttor, Florian Stimberg (TU Berlin)

Dan Cornford, Yuan Shen, Remi Barillec, Michail Vrettas (Aston U)

John Shawe–Taylor, Cédric Archambeau (UCL)

Guido Sanguinetti (Edinburgh U)

Overview

- Examples of partly observed continuous time Markov processes
- Reminder of Markov processes
- Monte Carlo, exact and linear noise approximations
- Variational approaches to inference
- Applications to a model of transcriptional regulation
- Outlook

Stochastic Lotka Volterra Model

2 interacting species (animals): predators and prey.

Numbers of animals: X_{Prey} and X_{Pred} .

$$X_{\text{Prey}} \rightarrow X_{\text{Prey}} + 1$$
 with Rate αX_{Prey}
 $X_{\text{Prey}} \rightarrow X_{\text{Prey}} - 1$ with Rate $\beta X_{\text{Prey}} X_{\text{Pred}}$
 $X_{\text{Pred}} \rightarrow X_{\text{Pred}} + 1$ with Rate $\delta X_{\text{Prey}} X_{\text{Pred}}$
 $X_{\text{Pred}} \rightarrow X_{\text{Pred}} - 1$ with Rate γX_{Pred}



A realisation of the stochastic process and the reaction constants



Discrete noisy observations from the continuous time series



6



Simple model of autoregulatory network

2 interacting molecules: mRNA and a Protein

Number of mRNA and Protein molecules: X_{mRNA} , X_{Prot}

 $\mathbf{X} = (X_{\mathsf{mRNA}}, X_{\mathsf{Prot}})$

$$\begin{array}{rcl} X_{\rm Prot} & \rightarrow & X_{\rm Prot} + 1 & : \ \text{with Rate} & \gamma X_{\rm mRNA} \,, \\ X_{\rm Prot} & \rightarrow & X_{\rm Prot} - 1 & : \ \text{with Rate} & \delta X_{\rm Prot} \,, \\ X_{\rm mRNA} & \rightarrow & X_{\rm mRNA} - 1 & : \ \text{with Rate} & \beta X_{\rm mRNA} \,, \\ X_{\rm mRNA} & \rightarrow & X_{\rm mRNA} + 1 & : \ \text{with Rate} & \alpha (1 - \alpha_c \Theta (X_{\rm Prot} - \theta_c)) \,, \end{array}$$

where $\Theta(x) = 1$ if $x \ge 0$ and $\Theta(x) = 0$ for x < 0.



Simulation of process



(Noisy) observations at discrete times.



Inference of state variable.

Motion in double-well potential

$$dX = f(X)dt + \sigma^2 dW.$$

with
$$f(x) = -\frac{dV(x)}{dx}$$

and V(x) is a double well potential



A sample path might look like this



Observations & optimal prediction



If we forget about the model and use a Bayesian regression ...



A reminder of Markov Processes

• Discrete time (Markov chain)

$$P(X_{n+1} \in A | X_n = x, X_{n-1} = x_{n-1}, \dots, X_0 = x_0)$$
$$= P(X_{n+1} \in A | X_n = x)$$

• Continuous time

$$P(X_{t+dt} \in A | X_{0:t} = x_{0:t}) = P(X_{t+dt} \in A | X_t = x_t)$$

where $x_{0:t} = \{x(s) | 0 \le s \le t\}$ denotes the entire path.

Jump Processes

Assume that X_t jumps between discrete states.

• Short time behaviour of transition kernel defined by **transition rate** *f*:

$$P(y,t + \Delta t|x,t) \simeq \delta_{xy} + f(y|x,t)\Delta t$$

for $\Delta t \to 0$. Note that $f(x|x,t) = -\sum_{z \neq x} f(z|x,t)$.

• Density for time t to next event (homogeneous process)

$$p(t) = -f(x|x) e^{f(x,x)t}$$

Discrete event simulation

- 1. initialise X_0 .
- 2. Let $X_t = x$. Simulate waiting time t' to next event with density $p(t') = -f(x|x) e^{f(x|x)t'}$.
- 3. Put t := t + t'
- 4. Draw new state x' with prob $-\frac{f(x'|x)}{f(x|x)}$
- 5. Output $X_t = x'$
- 6. If $t < T_{max}$ return to step 2.

This is the basis for the *Gillespie algorithm*.

Examples

Poisson Process: $X_t \in N_0$ and $f(x'|x) = \lambda \delta_{x',x+1}$



Telegraph Process: $X_t \in \{0, 1\}$

 $f(x'|x) = \lambda \delta_{x',1-x}$



Forward and Backward equations

• Forward (Master) equation

$$\frac{dP(x,t|x_0,t_0)}{dt} = \sum_{x' \neq x} \left[P(x',t|x_0,t_0) f(x|x') - P(x,t|x_0,t_0) f(x'|x) \right].$$

• Kolmogorov Backward equation

$$\frac{dP(x',t'|x,t)}{dt} = \sum_{y \neq x} f(y|x) \left\{ P(x',t'|x,t) - P(x',t'|y,t) \right\}$$

Diffusion approximation

Assume that jumps $\Delta x \doteq x' - x \ll x$. Try approximation by a continuous state Markov process

Introduce **drift** (1. Jump moment)

$$f(x) = \lim \frac{1}{\Delta t} E \left[X_{t+\Delta t} - X_t | X_t = x \right] = \sum_{\substack{x' \neq x}} f(x'|x)(x'-x),$$

and **diffusion matrix** (2. jump moment)

$$D(x) = \lim \frac{1}{\Delta t} E\left[(X_{t+\Delta t} - X_t)(X_{t+\Delta t} - X_t)^\top | X_t = x \right] = \sum_{\substack{x' \neq x}} (x' - x)f(x'|x)(x' - x)^\top$$

Drift & diffusion for Lotka - Volterra

$$h(X_1 \to X_1 + 1) = \alpha X_1, \quad h(X_1 \to X_1 - 1) = \beta X_1 X_2, h(X_2 \to X_2 + 1) = \delta X_1 X_2, \quad h(X_2 \to X_2 - 1) = \gamma X_2,$$

$$\mathbf{f}(\mathbf{x}) = \begin{pmatrix} \alpha x_1 - \beta x_1 x_2 \\ \delta x_1 x_2 - \gamma x_2 \end{pmatrix}$$
$$D(\mathbf{x}) = \begin{pmatrix} \alpha x_1 + \beta x_1 x_2 & 0 \\ 0 & \delta x_1 x_2 + \gamma x_2 \end{pmatrix}$$

Deterministic limit

Neglect fluctuations (diffusion), i.e. assume

$$X_{t+\Delta t} - X_t = f(X_t) \Delta t$$

Hence, $X_t \approx m(t) \doteq E[X_t]$. This yields 'classical' rate equation

 $\frac{dm}{dt} \approx f(m(t))$

Diffusion approximation

$$X_{t+\Delta t} - X_t = f(X_t)\Delta t + D^{1/2}(X_t)\sqrt{\Delta t} \epsilon_k$$

with ϵ_k i.i.d. ~ $\mathcal{N}(0, I)$, has the same 2 conditional moments f(x) and D(x). Jump sizes will go $\rightarrow 0$ and process becomes continuous!

We will write this (in the limit) as an **(Ito) stochastic differential** equation

$$dX_t = f(X_t)dt + D^{1/2}(X_t)dW_t$$

Short time behaviour of transition kernel

for Ito SDE

$$p\left(x',t+\Delta t|x,t\right) =$$

$$=\frac{1}{|2\pi D(x)|}\exp\left[-\frac{1}{2\Delta t}\left(\left(x'-x-f(x)\Delta t\right)^{\top}D^{-1}(x)\left(x'-x-f(x)\Delta t\right)\right]$$

Simulations of Jump and corresponding diffusion processes

protein



Forward and backward equations for diffusion processes

Let $x \doteq (x_1, ..., x_d)$ and $f(x) = (f_1(x), ..., f_d(x))$

• Forward (Fokker - Planck) equation

$$\frac{\partial p(x,t|x_0,t_0)}{\partial t} = -\sum_{i} \frac{\partial}{\partial x_i} (f_i(x)p(x,t|x_0,t_0)) + \frac{1}{2} \sum_{ij} \frac{\partial^2}{\partial x_i \partial x_j} \left(D_{ij}(x)p(x,t|x_0,t_0) \right)$$

• Kolmogorov Backward equation

$$\frac{\partial p(x',t'|x,t)}{\partial t} = -\sum_{i} f_i(x) \frac{\partial}{\partial x_i} p(x',t'|x,t) - \frac{1}{2} \sum_{ij} D_{ij}(x) \frac{\partial^2}{\partial x_i \partial x_j} p(x',t'|x,t)$$

Short summary

• Markov jump processes: Discrete states X

$$p\left(X', t + \Delta t | X, t\right) \simeq \delta_{X'X} + \Delta t f\left(X' | X\right)$$

for $\Delta t \rightarrow 0$

• Diffusion processes: (Ito) stochastic differential equation for state $X(t) \in \mathbb{R}^d$

$$dX(t) = \underbrace{f(X(t))}_{\text{Drift}} dt + \underbrace{D^{1/2}(X(t))}_{\text{Diffusion}} \times \underbrace{dW(t)}_{\text{Wiener process}}$$

Limit of discrete time process X_k

$$\Delta X_k \equiv X_{k+1} - X_k = f(X_k) \Delta t + D^{1/2}(X_k) \sqrt{\Delta t} \epsilon_k .$$

 ϵ_k i.i.d. Gaussian.

Inference Problems

Given noisy observations $D \equiv (y_1, \ldots, y_N)$ of hidden process $X(t_i)$ at times t_i for $i = 1, \ldots, N$.

- Estimate X(t) for $0 \le t \le T$ (**smoothing**).
- Estimate system parameters θ in drift f and diffusion D or in rates f(X'|X)

Can't we just discretise

and run an MCMC (Gibbs) sampler ?

- Alternate between sampling paths $p(x_{0:T}|\theta)$ and parameters $p(\theta|x_{0:T})$.
- Path sampling: Paths are described by a large vectors $x_{0:T} = (x_1, \dots x_K)$.

Sample from multivariate density

$$q(x_{0:T}) \propto e^{-V(x_{0:T})}$$

with

$$V(x_{0:T}) = \frac{1}{2\Delta t} \sum_{i} \left(x_{i+1} - x_i - f(x_i) \Delta t \right)^{\top} D^{-1}(x_i) \left(x_{i+1} - x_i - f(x_i) \Delta t \right)$$
$$-\sum_{i} \ln p(y_{obs} | x_{obs})$$

• Block updates using bridge proposals within Metropolis - Hastings algorithm (Golightly, Wilkinson 2008).



• Alternative: Hybrid Monte Carlo (Alexander et al 2005)

Problem with Gibbs samplers

• Problem: For $\Delta t \rightarrow 0$ (complete SDE path) diffusion D and $x_{0:T}$ are highly dependent:

$$\sum_{t} \left(X_{t+\Delta t} - X_t \right)^2 \to \int_0^T D(X_t) dt$$

• A Gibbs sampler will get stuck in the same (diffusion) parameters for small Δt .



Ornstein Uhlenbeck - processes for $\sigma = 1$ (red), $\sigma = 5$ (green), $\sigma = 10$ (blue).

Isn't this just some sort of hidden Markov model ?

- Yes, but we don't know the transition probabilities.
- Can't we still use a Forward Backward Algorithm ?

Forward - Backward Solution

• Marginal posterior at time t

$$p_t(x|\{y_i\}_{i=1}^N) \propto \underbrace{p(x_t|\{y_i\}_{t_i < t})}_{\doteq p_F(x,t)} \times \underbrace{p(\{y_i\}_{t_i \ge t}|X_t = x)}_{\doteq \psi(x,t)}$$

• Fokker - Planck equation (between observations)

$$\left\{\frac{\partial}{\partial t} + \nabla f - \frac{1}{2} \operatorname{Tr}(\nabla \nabla^T D)\right\} p_F(x,t) = 0$$

• Backward equation

$$\left\{\frac{\partial}{\partial t} + f^{\top} \nabla + \frac{1}{2} \operatorname{Tr}(D \nabla^{\top} \nabla)\right\} \psi(x, t) = 0$$

• & a set of jump conditions at observations.

KSP equations (Kushner '62, Stratonovich '60 & Pardoux '82).

The backward equation is important !

• Posterior process is Markovian and fulfils SDE

$$dX(t) = g(X(t), t)dt + D^{1/2}(X(t)) dW(t)$$

with an effective time dependent drift g(X(t), t)

$$g(x,t) = f(x) + D(x,t)\nabla \ln \psi(x,t)$$

and ψ is the solution to the backward equation.

• Likelihood of observations $P(\mathbf{y}_1, \dots, \mathbf{y}_n | X_0 = x) = \psi(x, 0)$.
Example

Wiener process with single, noise free observation y = x(t = T) = 0



Posterior drift $g(x,t) = -\frac{x}{T-t}$ for 0 < t < T.

The linear noise approximation

- Similar to extended Kalman filter approach.
- Expansion in small parameter like 1/Number of molecules (Van Kampen).
- leads to linear SDE: X_t becomes approximated by a Gaussian process !
- Can be applied to backward equation \rightarrow ODEs instead of PDEs !

Comparison of linear noise with MCMC





Posterior for parameters



Comments

- Good approximation for small fluctuations
- Does not work well for multistable systems
- Gaussian approximation not optimised !

The variational approach in machine learning

Let
$$P(X|D) = \frac{P(D|X)P(X)}{P(D)}$$
. Then we have
 $P(\cdot|D) = \arg\min_{Q} \mathcal{F}[Q]$

where the variational free energy is defined as

$$\mathcal{F}[Q] = KL(Q|P_{prior}) - E_Q[\ln P(D|\cdot)]$$

with the Kullback - Leibler divergence

$$KL(Q|P) = E_Q[\ln\frac{Q}{P}]$$

٠

and

$$\min_{Q} \mathcal{F}[Q] = P(D)$$

Parameter inference $P(D) \rightarrow p(D|\theta)$: Approximate posterior of parameters (Lappalainen, 2000)

$$q(\theta|D) \approx \frac{e^{-\mathcal{F}_{\theta}[q]} p(\theta)}{\int e^{-\mathcal{F}_{\theta}[q]} p(\theta) d\theta}.$$

The posterior measure

Conditional (posterior) distribution over **paths** $X_{0:T}$ given the data $D = (y_1, \ldots, y_n)$

$$P(X_{0:T}|D) = \frac{P_{prior}(X_{0:T})}{P(D)} \times \prod_{n=1}^{N} p(y_n|X(t_n)),$$

The Kullback - Leibler divergence for paths

Discretize !

Prior and posterior over paths are Markovian with transition probabilities $P(x', t'|x, t_k)$ and $Q(x', t'|x, t_k)$.

$$KL[Q||P] = \int dx_{0:T} Q(x_{0:T}) \ln \frac{Q(x_{0:T})}{P(x_{0:T})}$$

= $\sum_{k=0}^{K-1} \int dx Q(x,t_k) \int dx' Q(x',t_k + \Delta t|x,t_k) \ln \frac{Q(x',t_k + \Delta t|x,t_k)}{P(x',t_k + \Delta t|x,t_k)}$
= $\sum_{k=0}^{K-1} \int dx Q(x,t_k) KL[Q(\cdot,t_k + \Delta t|x,t_k) ||P(\cdot,t_k + \Delta t|x,t_k)]$

... use short time transition probability

For diffusions

$$p\left(X', t + \Delta t | X, t\right) \propto \exp\left[-\frac{1}{2\Delta t} \|\Delta X - f(X)\Delta t\|_D^2\right]$$

as $\Delta t \to 0$,

with $||F||_D^2 = F^\top D^{-1} F$.

and take the limit $\Delta t \rightarrow 0$

Let Q and P be measures over paths for SDEs with drifts g(X,t) and f(X,t) with same diffusion D(X). Then

$$KL[Q||P] = \frac{1}{2} \int_0^T dt \left\{ \int dx \ q(x,t) \ ||g(x,t) - f(x,t)||_D^2 \right\}$$

q(x,t) is the density of X(t).

The variational problem (exact inference !)

Minimise variational free energy

$$\mathcal{F}_{\theta}(Q) = \frac{1}{2} \int_0^T \int q(x,t) \, \|g(x,t) - f(x)\|_D^2 \, dx \, dt \, -\sum_i E_Q[\ln p(y_i|X(t_i))]$$

The marginal density q fulfils the Fokker - Planck equation

$$\frac{\partial q}{\partial t} = \left\{ -\nabla g + \frac{1}{2} \operatorname{Tr}(\nabla \nabla^T) D \right\} q \equiv L_g q$$

Variation of Lagrange function

$$L = \frac{1}{2} \int_0^T dt \int dx \, \|g(x,t) - f(x,t)\|_D^2 \, q(x,t)$$
$$-\sum_i \int dx \, q(x,t_i) \ln p(y_i|x)]$$
$$-\int_0^T dt \int dx \, \lambda(x,t) \, \left(\frac{\partial q}{\partial t} - L_g \, q\right)$$

with respect to q and g, setting $\lambda(x,t) = -\ln \psi(x,t)$ yields

$$g(x,t) = f(x) + D(x,t)\nabla \ln \psi(x,t)$$
$$\left\{\frac{\partial}{\partial t} + f^{\top}\nabla + \frac{1}{2}\mathrm{Tr}(D\nabla^{\top}\nabla)\right\}\psi(x,t) = 0$$

+ jump conditions at observations.

Relation to optimal control

(B Kappen, E Todorov)

Consider SDE

$$dX(t) = (u(X(t), t) + f(X(t)))dt + D^{1/2}(X(t)) dW(t)$$

Adjust control $u(X_t, t)$ so that

$$V = \int_0^T E^u \left[L_t(X_t, u_t) \right]$$

is minimal, where

$$L_t(x, u) = \frac{1}{2} \|u\|_D^2 + U(x(t), t)$$

The Gaussian Variational Approximation



For previous applications in machine learning (see e.g. Barber & Bishop (1998), Seeger (2000), Honkela & Valpola (2005)).

The Gaussian Variational Approximation for SDEs

(Archambeau, Cornford, Opper & Shawe - Taylor, 2007)

- Diffusion must be independent of X.
- Gaussian measure over paths $X_{0:T}$ induced by linear (approximate) posterior SDE:

$$dX(t) = \{-A(t)X + b(t)\} dt + D^{1/2} dW$$

Constraints are evolution eqs. for marginal mean and covariance

$$\frac{dm}{dt} = -Am + b(t)$$
$$\frac{dS}{dt} = -AS - SA^{\top} + D$$

 \rightarrow nonlinear ODEs instead of PDEs !

Motion in double-well potential

 $dX = 4X(1 - X^2)dt + \sigma^2 dW.$



Variational result and comparison to MCMC

(Archambeau, Opper, Shen, Cornford, Shawe-Taylor, 2008)



Estimation of Diffusion constant σ^2

$$q(\theta|D) \approx \frac{e^{-\mathcal{F}_{\theta}(Q)} p(\theta)}{\int e^{-\mathcal{F}_{\theta}(Q)} p(\theta) d\theta}$$



With more observations



More dimensions

Toy weather model (Lorenz and Emanuel, 1998): $x = (x^1, \dots, x^{40})$ with drift



$$(x_t) = \left(x_t^{i+1} - x_t^{i-2}\right) x_t^{i-1} - x_t^i + \theta$$

(Vrettas, Cornford and Opper, submitted)

Large observation noise



Double well with observation noise $\sigma_o = 0.6$

KL divergence for Markov jump processes

Assume transition rates g(X'|X,t) and f(X'|X,t)

$$KL[Q||P] = \int_0^T dt \sum_x q(x,t) \sum_{x':x' \neq x} \left\{ g(x'|x,t) \ln \frac{g(x'|x,t)}{f(x'|x,t)} + f(x'|x,t) - g(x'|x,t) \right\}$$

Mean field approximation

Multivariate states $X = (X(1), \ldots, X(d))$

Exact inference: Linear ODEs in S^d variables

Variational approximation: Optimise in family of factorising measures, i.e. of the type

 $Q[X_{0:T}] = \prod_{i=1}^{d} Q_i[X_{0:T}(i)]$

Linear ODEs in Sd variables.

(Sanguinetti & Opper, 2008)

Simple model of autoregulatory network

2 interacting molecules: mRNA and a Protein

Number of mRNA and Protein molecules: X_{mRNA} , X_{Prot}

 $\mathbf{X} = (X_{\mathsf{mRNA}}, X_{\mathsf{Prot}})$

$$\begin{array}{rcl} X_{\rm Prot} & \rightarrow & X_{\rm Prot} + 1 & : \ \text{with Rate} & \gamma X_{\rm mRNA} \,, \\ X_{\rm Prot} & \rightarrow & X_{\rm Prot} - 1 & : \ \text{with Rate} & \delta X_{\rm Prot} \,, \\ X_{\rm mRNA} & \rightarrow & X_{\rm mRNA} - 1 & : \ \text{with Rate} & \beta X_{\rm mRNA} \,, \\ X_{\rm mRNA} & \rightarrow & X_{\rm mRNA} + 1 & : \ \text{with Rate} & \alpha (1 - \alpha_c \Theta (X_{\rm Prot} - \theta_c)) \,, \end{array}$$

where $\Theta(x) = 1$ if $x \ge 0$ and $\Theta(x) = 0$ for x < 0.



(M. Opper & G. Sanguinetti, NIPS07)

Critical parameter $\theta_c = 20$ (exact) and $\theta_c = 19$ (ML estimate).



- Transcription factors regulate genes by binding to specific sites.
- Hard to measure transcription factor activity directly. Inference must be based on measurement of mRNA concentration of target genes.
- Big networks: Use Clustering. Groups of genes with similar expression profiles likely regulated by same TF.
- Big networks: Latent variable models (Factor analysis).

• Small subnetworks: more detailed dynamical model (Barenco et al:)

$$\frac{dx_i}{dt} = -\lambda_i x_i(t) + A_i f(t)$$

takes sensitivity and degradation into account.

Try predictions on TF activity f(t) and learn parameters using measurements of mRNA concentration of target genes:

$$y_{ik} = x_i(t_k) + \text{noise}$$

A priori knowledge of transcription factor activity f(t)

- Discretize and assume f(t) independent for different times t (Barenco et al, 2006).
- f(t) smooth & modelled by a Gaussian process (Lawrence, Sanguinetti & Rattray, 2006)
- This work: Fast switching activity $\mu(t) \in \{0, 1\}$ and $\mu \to 1 \mu$ with rates γ_{\pm} . (Sanguinetti, Ruttor, Archambeau, Opper 2009)
- The latter model is

$$\frac{dx_i}{dt} = -\lambda_i x_i(t) + A_i \mu(t) + b_i$$

Telegraph Process: $X_t \in \{0, 1\}$





The role of the prior: telegraph vs GP

Variational approximation for Telegraph process

• We need the *variational free energy*

$$\mathcal{F}[Q] = KL(Q|P_{prior}) - E_Q[\ln p(D|\cdot)]$$

where Q is a Markov process over binary paths $\mu_{0:T}$.

- $D = (y_1, \ldots, y_n)$ are noisy observations of $x(t_i)$
- where x(t) is a solution to

$$\frac{dx_i}{dt} = -\lambda_i x_i(t) + A_i \mu(t) + b_i$$

• Hence

$$p(D|\mu_{0:T}) \propto \exp\left[-\frac{1}{2\sigma^2}\sum_{i=1}^n (x(t_i) - y_i)^2\right]$$

• the linear ODE is solved by

$$x(t) = \exp(-\lambda t) \left\{ x(0) + \int_0^t \exp(\lambda s) \left[A\mu(s) + b\right] ds \right\}$$

• Let f_{\pm} be the prior rates and $g_{\pm}(t)$ the posterior rates. Then

$$KL\left[Q\|P_{prior}\right] = \int_{0}^{T} dt \ q_{1}(t) \left[g_{-}(t) \ln \frac{g_{-}(t)}{f_{-}(t)} + f_{-}(t) - g_{-}(t)\right] + \int_{0}^{T} dt \ \left[1 - q_{1}(t)\right] \left[g_{+}(t) \ln \frac{g_{+}(t)}{f_{+}(t)} + f_{+}(t) - g_{+}(t)\right]$$

The total free energy

$$\mathcal{F}[q] = \int_0^T dt \ q_1(t) \left[g_-(t) \ln \frac{g_-(t)}{f_-(t)} + f_-(t) - g_-(t) \right] + \int_0^T dt \ \left[1 - q_1(t) \right] \left[g_+(t) \ln \frac{g_+(t)}{f_+(t)} + f_+(t) - g_+(t) \right] + \frac{1}{2\sigma^2} \sum_{i=1}^n E_q \left(\int_0^t \exp\left(-\lambda(t-s) \left[A\mu\left(s\right) + b\right] ds - y_i \right)^2 \right]$$

has to be minimised wrt $g_{\pm}(t)$ and $q_1(t)$ under the condition that

$$\frac{dq_1(t)}{dt} = -(g_+ + g_-)q_1(t) + g_+(t)$$



Approximate vs exact inference

62


based on measurements on 5 target genes at 4 time points.

(Sanguinetti et al, 2009)

Multiple (2) transcription factors (Opper & Sanguinetti, 2010):

$$\frac{dx_i}{dt} = -\lambda_i x_i(t) + A^i \mu_1(t) + A^i \mu_1(t) + A^i_{12} \mu_1(t) \mu_2(t) + b_i$$





Parameter inference for A_1^2 (blue), A_2^2 (green) and A_{12}^2 (red).



Prediction of activity of transcription factors **FHL1** and **RAP1** (Microarray data from yeast metabolic cycle)







 A_1^i , A_2^i and A_{12}^i for 5 target genes.

Comments

- Method scales linearly in number of genes, number of observations & number of TFs.
- We can calculate upper and lower bounds on free energy.
- Can be generalised to more complex architectures (forward loops, autoregulatory networks).
- Can be generalised to additional white noise input (linear SDEs).

Inclusion of white noise (for single cell fluorescence measurements)

$$dx_i = (-\lambda_i x_i(t) + A_i \mu(t) + b_i) dt + \sigma dW$$

Toy data



Opper, Ruttor and Sanguinetti, NIPS2010

Present & Future work

- Variational path densities as proposal for MCMC
- Perturbative corrections, PAC Bayesian confidence bounds
- Variational approaches to multiplicative noise ?
- Realistic higher dimensional systems: Need good suboptimal variational parametrisations.